Department of Mathematics Faculty of Mathematics and Natural Sciences, Andalas University, Indonesia

The 5th IMT - GT **International Conference** on Mathematics, Statistics and Their Applications

Editors: I Made Arnawa, Muhafzan, Maiyastri, Susila Bahri

ICMSA 2009





Organized by:

June 9-11, 2009 The Hills Hotel Bukittinggi, Indonesia











Proceedings of

The 5th IMT-GT

International Conference on Mathematics, Statistics, and their Applications 2009

(ICMSA 2009)

"Mathematics and Statistics for Industry and Community Development"

June 09 - 11, 2009

The Hills Hotel, Bukittinggi - Indonesia

Organized by:
Department of Mathematics, Andalas University,
Padang – Indonesia

Published by:

Department of Mathematics, Andalas University Kampus UNAND Limau Manis Padang 25163, Sumatera Barat, INDONESIA.

ISBN 978-602-95343-0-6

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June 2009

Preface

First of all, I would like to say welcome to Bukittinggi, Indonesia to all of you. It is an honour for us to host this conference. We are very happy and proud because the participants of this conference come from many countries; we have participants from Libya, Japan, Qatar, India, Malaysia, Singapore, Thailand, Iran, and many more.

Ladies and gentlemen, according to constructivism theory, mathematics comes out as a result of social construction; that's why, the outcomes of our researches in mathematics, like theorem or formula of mathematics, should be communicated in a scientific forum such as seminar or conference. Through this kind of seminar or conference, we could refine the existing theorems or we could get new ideas to produce a new one. Seminar or conference which is held annually enables us to continually develop the science of mathematics until the end of the time.

That's way, in this two-day conference, we are going to discuss around 250 papers coming from diverse aspects of mathematics ranging from analysis, applied mathematics, statistics, algebra, Computational Mathematics, mathematics education, and other related topics.

For all of us here, I would like to convey my endless appreciation and gratitude for your participation in this conference.

Thank you very much

Dr. I Made Arnawa

Chairman of the Conference

Message from Rector Andalas University

It gives me great pleasure to extend my sincere and warm welcome to the participants of the 5th International Conference on Mathematics Statistics and Application (The IMT GT's 5th ICMSA 2009) - A Joint Scientific Program organized by Universities over Indonesia, Malaysia and Thailand Growth Triangle Region. On behalf of Andalas University, let me welcome all of you to the conference in Bukittinggi, West Sumatra Province, the land of Minang kabau.

We believe that from this scientific meeting, all of participants will have time to discuss and exchange ideas, findings, creating new networking as well as strengthen the existing collaboration in the respective fields of expertise. In the century in which the information is spreading in a tremendous speed and globalization is a trend, Andalas University must prepare for the tough competition that lay a head. One way to succeed is by initiating and developing collaborative work with many partners from all over the world. Through the collaboration in this conference we can improve the quality of our researches as well as teaching and learning process in mathematics and to achieve standards and requirements applied in many developed countries. I strongly believe that this conference is and extraordinary testimony to our capacity building at international, regional and local level.

I would like to express my deep gratitude to International Scientific Committee of who has honored the Mathematics Department, Faculty of Mathematics and Natural Sciences, Andalas University to host this prestigious conference. This is a very special opportunity for us to be involved in a regional community of knowledgeable scientist in the field of mathematics, statistics and their applications. I would also like to extend my gratitude to keynote speakers, participants, and organizer of this conference for their contribution to this event. My special thank is also rendered to the local government of West Sumatra for various supports and facilities.

Finally I wish all participants a fruitful deliberation at the conference. I also wish all participants and accompanying spouses a pleasant and enjoyable stay in Bukittinggi City, West Sumatra.

Prof. Dr. Ir. Musliar Kasim, MS

Lune has

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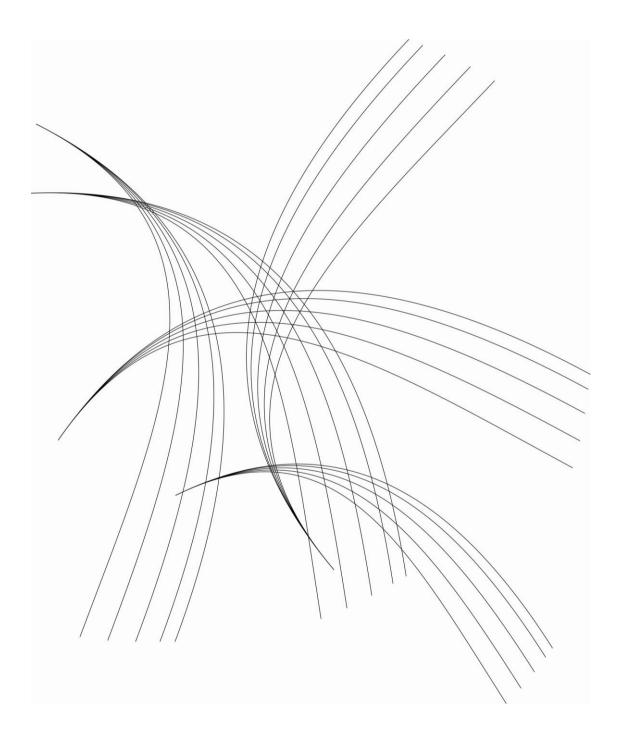
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Contents



CONTENTS

Preface	1
Message from Rector of Andalas University	ii
Organizing Committee	iii
Algebra	1 - 54
A subclass of 7-dimensional complex filiform Leibniz algebras and their isomorphism classes Sharifah Kartini Said Husain, Isamiddin S. Rakhimov	1
New key exchange in Elliptic Curve based on Decomposition Problem Hilyati Hanina Zazali, Wan Ainun Mior Othman	7
Generators of Diagram Groups From Semigroup Presentation $P = \langle x, y, z x = y, y = z, x = z \rangle$ Using Lifting Methods Yousof Gheisari and Abd Ghafur Bin Ahmad	11
Finite Field Basis Conversion Intan Muchtadi-Alamsyah, Marisa W. Paryasto, Muhammad Hafiz Khusyairi	15
On Commutative Group of N-homomorphisms Indah Emilia Wijayanti	19
The <i>f</i> -chromatic indexes of wheel-like graphs <i>Adiwijaya</i> ^{1,2} , <i>A.N.M. Salman</i> ¹ , <i>D. Suprijanto</i> ¹ , <i>E.T. Baskoro</i> ¹	24
Structure of Incidence Algebras of Locally Finite Partially Ordered Set Gustina Efiyanti ¹ and Irawati ²	28
A new directed hypergraph distance and its applications Mulia Astuti ¹ , Irawati ² , Intan Muchtadi-Alamsyah, Ahmad Muchlis ⁴ , Achirul Akbar ⁵ dan Muliana. A. Halim ⁶	31
Large graphs diameter two from the smaller ones Yus Mochamad Cholily	37
On Module Classes Closed Under Submodules, Factor Modules and Direct Sums Suwarno Ariswoyo and Elvina Herawaty	40
Multinomial Option Pricing with Pseudoinverse Matrix Abdurakhman	44
Solving Traveling Salesman Problem Using A Hibryd of Evolution Strategies and Lin-Kernighan Algorithm Nurmaulidar	46
Analysis	55 - 84
Generation of Rainfall Sequence using Fourier Series Norzaida Abas ^{1*} , Zalina Mohd Daud ¹ , Fadhilah Yusof ²	55
On Extremal Properties for Certain Classes of Analytic Functions Shaharuddin Cik Soh and Daud Mohamad	61
The Area of the Region Enclosed by Bezier Curve Normi binti Abdul Hadi*, ² Abdul Halmie bin Muhamad	66
The Application of Regulated Function on the Multiplication of Two Henstock Integrable Function Christiana Rini Indrati	70
Riesz potential and the generalized Morrey spaces with growth measures Idha Sihwaningrum ^{1*} , Hendra Gunawan ² , Yudi Soeharyadi ³ , Wono Setya Budhi ⁴	75
An Estimation of Exponential Sums Associated With A Sextic Form Sapar S. H^{l} , Mohd Atan K. A^{2}	79

Computational Mathematics	85 - 180
A Zero-dissipative Runge-Kutta-Nyström Method with Minimal Phase-lag for Oscillatory Problems Norazak Senu ^{1*} , Mohamed Suleiman ² Fudziah Ismail ³ and Mohamed Othman ⁴	85
New Multi-step Runge-Kutta method O. Y. Ababneh ^{1*} , R. Ahmad ² , E. S. Ismail ³	91
An efficient parallel implementation of Markov clustering algorithm for large-scale protein-protein interaction networks that uses MPI Alhadi Bustamam ^{1, 2, *} , Muhammad Shoaib Sehgal ¹ , Nicholas Hamilton ¹ , Simon Wong ¹ , Mark A Ragan ¹ , Kevin Burrage ^{1, 3}	94
Implementation Of Parallel Computational Tools For The Curing Simulation Of Thermoset Composites Using The Two Dimension Age Algorithm Amna Abdurrahman ¹ , Ahmad Kamal bin Zulkifle ² , Norma Alias ³ , and Ishak Hashim ⁴	102
Optimizing of Text Retrieval: A similarity level by keyword competition in Genetic Algorithm (GA) $Poltak\ Sihombing^I$	109
The mutation and crossover effect in genetic algorithm to determine the similarity level of document retrieval <i>Poltak Sihombing</i> ¹	117
Analysis and Evaluation RC4 Algorithm for Data Encryption Ratna Aisuwarya ¹ , Rahmi Eka Putri ²	124
Efficient Differential Equation Solvers for Fluid Modeling in Interactive Surgical Drilling Simulations based on GPU Computations Sugeng Rianto ^{1,*} Ling Li ²	132
Developing Secant Method for Solving Nonlinear Equations Taufiq Iskandar ¹ , Marlan ² , Sanggam P. Gultom ³ , Herman Mawengkang ⁴	137
Visualization of XLM-based Geochemical Data using SVG and ASP.NET $Nizamuddin^{I}*$ and $Hidehiro\ Ishizuka^{2}$	151
Application of Backtracking Algorithm On the Knight's Tour Game Using Pascal Senja Omega Puspita, Budi Rudianto	156
A Mixed Integer Linear Programming Model for Capacity Selection Problem in Logistics Networks Optimization <i>Muhammad Izman Herdiansyah</i> ¹	164
Analysis and Implementation of the New Student Acceptance System On 2 ^{sd} Padang's Educational Council in On-Line Method By Using Applserver 2.5.7 Application Mohammad Hafiz Hersyah	175
Statistic	181 - 389
Goodness of Fit Test for the EEG Distribution by Using the Empirical Laplace Transform Prasong Kitidamrongsuk ¹ *, Pachitjanut Siripanich ²	181
Conditional Maximum Likelihood Estimator for Incomplete Longitudinal Data Juthaphorn Saekhoo ¹ *, Pachitjanut Siripanich ²	186
Autocorrelation correction in a simultaneous equations model <i>Kerativibooly W.</i> ^{1*} , <i>Jitthavech J., Lorchirachoonkul V.</i>	192
Construction of Weights for the Estimation of Regression Coefficients with Outliers <i>Pimpan Ampanthong</i> ¹ , <i>Prachoom Suwattee</i> ²	198
Goodness of Fit for the Poisson Distribution Based on Sample Skewness Manad Khamkong* and Pachitjanut Siripanich	204
Estimators of Errors of Kernel Density Estimator Manachai Rodchuen 1*, Prachoom Suwattee 1	209

Using Data Envelopment Analysis (DEA) Model to Measure The Relative Efficiency Of Higher Educations	ion
¹ Zalina Zahid, ¹ Rasimah Aripin and ² Mohd. Nasir Taib	216
Detecting Students at Risk of Failing Using Decision Tree Haliza Hasan*, Rasimah Aripin & Sharifah Sakinah Syed Hassan Aidid	222
Assessing The Performance of University Departments Using Data Envelopment Analysis Nordin Hj Mohamad ^{1*} and Fatimah Said ²	228
Mean-VaR Portfolio Optimization Under CAPM by Non Constant Volatility in Return Market <i>Sukono¹</i> , <i>Subanar²</i> & <i>Dedi Rosadi³</i>	238
Credit Risk Measurement for a Single Facility in Banking Sukono	242
Asymptotic Properties of a Generalized Renewal Reward Process Suyono	248
Distribution of a Generalized Renewal Reward Process Suyono ¹⁾ and Subanar ²⁾	252
Bayesian Updating Reservoir Simulation Models Sutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Agus Yodi Gunawan ² , Sri Wahyuningsih ¹ , Nurtiti Sunusi ¹ , Aceng Komarudin Mutawanir Darwis ^{1*} , Aceng	256 qin ¹
Developing Bayesian inferential method for basic reproduction number in epidemic models under	r
uncertainty Dapot Situngkir, Abdul Latif Hasibuan, Daswati Sigalingging, Debora S.Parapat, Erwin Sidabal Herman Mawengkang	lok, 261
Two stage mixed integer nonlinear stochastic programming model for solving a superstructure synthesis water networks optimization problem under uncertainty Mujio, Evi Yanti Lubis, Harris H.Simamora, Herbin Manurung, Januasi Simarmata, Herman Mawengka	270
Application of Cluster Analysis In Classification of Tourist Destinations In Sabang ¹ Asep Rusyana, ² Evi Ramadhani, ³ Suhartono	278
Simulation of Agent Based Model on Jakarta Stock Exchange (BEJ) Afdal Mazni	281
Modeling and Analyzing Dependent Categorical Data Georgina M. Tinungki	296
Life Insurance Product Valuation under Binomial Framework Danang Teguh Qoyyimi*, Danardono, Abdurakhman¹	304
$\label{thm:covariates} \mbox{Time-dependent Strata as an alternative for Time-dependent Covariates in Survival Analysis $Danardono^{l}$$	313
Statistical Analysis on Determination of Optimum Condition for Pra-esterification of Crude Palm Oil with Methanol to Biodiesel Using Sulfuric Acid Catalyst	
Sawaluddin ¹ , Tirena Bahnur Siregar ² , Suwarno Ariswoyo ³	318
A Linear Mixed Model for Two-Dimensional Competition between Neighbouring Trees in Forestry Tria Model Testing in a Plantation of Maritime Pine (<i>Pinus pinaster</i> Ait.) in Western Australia <i>Dadan Kusnandar</i> ^{1,*} , N.W. Galwey ²	als: 322
Reducing Fuzzy Relations of Fuzzy Time Series Model Using <i>QR</i> Factorization Method and Its Applica	ation
to Forecasting Interest Rate of Bank Indonesia Certificate Agus Maman Abadi ¹ , Subanar ² , Widodo ³ , Samsubar Saleh ⁴	328
An Adaptive Sensitivity-Based Linear Learning Method Algorithm for Data Classification <i>Zaenal Arifin</i> ¹⁾ , <i>M. Isa Irawan</i> ²⁾	333
The Study of Sensitivity of Radial Basis Probabilistic Neural Network <i>Hasanuddin</i> ^{1*} , <i>M. Isa Irawan</i> ²	344
Recurrence Time Modeling for Earthquake Prediction Surianto ¹ , Sutawanir Darwis ² , Aceng Komaruddin Mutaqin ³	350

Stochastic Project Scheduling Johannes P. Sitanggang, Budi Irwansyah, Risna Helvida, Yuliani Nasution, Syafaruddin, Herman Mawengkang	352
Mewma And Ewma Quality Control Charts, And Its Multiobjective Optimization Arrival Rince Putri ¹	364
Characteristic of Poisson Distribution Nova Noliza Bakar	368
The Comparison of Forecasting of Jakarta Composite Index by Using Exponential Smoothing and Arima	a
Method Puspa Amelia ¹⁾ and Maiyastri ²⁾	374
Representing Students Perspectives to Television Channels by Using Multidimensional Scaling Method (Case study in Department of Mathematics Andalas University) *Primawati**, Izzati Rahmi, HG*	381
Latin Square Arrangement for Taste-Panel Experiments Arisman Adnan	386
Mathematics Education 390	- 444
Improving Student Academic Performance by An Application of Data Mining Techniques Sajadin Sembiring ^{1*} , Abdullah Embong ² , Mohd. Azwan Mohamad ³ , Muhammad Furqan ⁴ .	390
To consider the education of mathematics in different educational levels and it's more effective learning methods Manochehr Kazemi ¹ and Hassan Naraghi ²	395
How to Make the Teaching of Abstract Algebra Interesting? Abdul Razak Salleh	397
Matriculation Students' Metacognitive Awareness And Achievement In Mathematical Problem Solving $\it Effandi\ Zakaria^1*$, $\it Zainah\ Yazid^2$	403
How The Indian Vedic Mathematics on Method of Calculation is still Relevant for Children in Today's Era of Mental Computation <i>Rita Desfitri</i> ^{1 2}	409
Prioritization of Factors to Further Studies in a University Using AHP Yuzainee Bte Md Yusoff ¹ , Norngainy Bt Mohd Tawil ²	416
Contribution of Games to the Student Interest on Mathematics Lesson <i>Syukma Netti</i> ¹ , <i>Niniwati</i> ²	422
Implementation of Creative Problem Solving Method by Using Media Computer to Improve Students' Achievement in Mathematics (Classroom Action Research at SMP 13 Pekanbaru) ¹ Yenita Roza, ² Maida Deli	424
Mathematic for Senior High School It's Problems and Challenges <i>Ali Asmar</i>	428
Errors in College Level Theorem Proving I Made Arnawa	431
Development of Model - Mathematics Learning Based of Interpersonal Intelligences For Student Class VII in Padang Atus Amadi Putra	436
Teaching Mathematics Through Cooperative Learning and Using ICT Hendra Syarifuddin	440
Improving Students' Activities And Mathematics Achievement Of SMPN 26 Padang Through Cooperative Learning; STAD Type <i>Mirna</i>	442

Applied Mathematics	445 - 979
The modification of steepest descent method On portfolio selection <i>Yosza Dasril</i> ¹ , <i>Goh Khang Wen</i> ² , & <i>Ismail Bin Mohd</i> ³	445
Feedback Control and Magnetic Field Effects on Marangoni Instability in a Micropolar Fluid Mohd Nasir Mahmud ¹ , Zainol Mustafa ² , and Ishak Hashim ^{2*}	451
A Mathematical Models of Lower Limb Using Kane's Method: An Application to Walking Movement While Carrying Load Fazrolrozi & A. S. Rambely	457
Mixed convection boundary layer flow towards a vertical plate embedded in a porous medium Norfifah Bachok ^{1*} , Anuar Ishak ² , Roslinda Nazar ³ , Ioan Pop ⁴	464
A New Number Representation for Faster Elliptic Curve Scalar Multiplication Mohamad Rushdan Md Said*, Abdul Wahed M. Ismail	471
Effect of Non-Uniform Temperature Gradient and Magnetic Field on Marangoni Convection in a Mic Fluid	cropolar
^{1,*} Melviana Johnson Fu, ² Norihan Md. Arifin, ³ Mohd Noor Saad, ⁴ Roslinda Mohd Nazar	476
MHD stagnation-point flow towards a stretching sheet with induced magnetic field Fadzilah Md Ali ^{1*} , Roslinda Mohd Nazar ² , Norihan Md Arifin ¹ , Ioan Pop ³	482
Multiple Intelligence Profiling Analysis of People with Epilepsy for Job Placement Purposes Siti Rahmah Awang ^{1*} , Rasimah Aripin ² , Md. Hanip Rafia ³	488
Mathematical Modeling of the Upper Limb Movement using Kane's Equation Sharifah Alwiah AbdulRahman 1*, Azmin Sham Rambely², Rokiah Rozita Ahmad³	495
Pixel area validation of segmented malignant tumors in digital mammographic images Rohana Embong ^{1*} , Wan Eny Zarina W.Abd. Rahman ¹ , Tahir Ahmad ² , Rozi Mahmud ³ , Arsmah Ibrahim ¹ , Zainab Abu Bakar ¹ , Md Saion Salikin ⁴	499
Structural Similarity Measure for Mathematics Assessment Marking Engine ¹ Arsmah Ibrahim, ² Zainab Abu Bakar, ³ Nuru'l – 'Izzah Othman	503
A Decision Making Model Based on Consistent Fuzzy Preference Relations and Generalized TOPSIS Nor Hanimah Kamis, Daud Mohamad and Nor Hashimah Sulaiman	S 509
The Study of Delivery Response Time: A Gap Analysis Approach ¹ Mohd Sahar Sauian, ² Mohamad Yazid Shuaib	516
Mathematical Musyarakah Model in Managing Islamic Investment Between Two Parties Using Two	Profit
Sharing Rates Maheran Mohd Jaffar	520
A Fuzzy Ruled Based Model for Stock Selection Advisor System Daud Mohamad, Noorhar Jiana Haryanti Mohd Saad	525
Performance of A Low-Cost Pcs in Edge Detection of Breast Tumor in Digital Mammomograms Usi Wavelet Modulus Maxima Arsmah Ibrahim ¹ , Norma Alias ² , Hanifah Sulaiman ¹ , Mohd Idris Jayes ¹ , Khairil Iskandar Othman ¹ , Saion Salikin ³	_
A Metaheuristics Approach for the Inventory Routing problem Huda Zuhrah Ab Halim*, Nur Arina Bazilah Aziz, Noor Hasnah Moin	535
Single-vendor single-buyer model under linearly decreasing demand Supadi, S.S ^{1*} , Omar, M. ²	540
Catastrophe Reinsurance Somayeh Nik Manesh 1*, Dr.Noor Azlinna Azizan 2	546
Determination of Work Done by a Female Student While Carrying Backpack with Different Loads ¹ Nor Atikah Ab Ghani and ² Azmin Sham Rambely	556

Linearity in Between House Price and Annual Management Fund in High-Rise Residential in Kuala Lun	npur,
Malaysia N. M. Tawil ¹ ., M. N.Daud ³ , A.I. Che-Ani ² , N.A.Goh ² , M.F.M.Zain ²	561
Fuzzy Conjoint Analysis of Influence Factors in High-Rise Residential Price And Management Fund In	
Kuala Lumpur, Malaysia N. Mohd-Tawil ¹ , A.I.Che-Ani ² , Amiruddin Ismail ³ , M.M.Tahir ² , M. Jamil ²	568
Value-based Total Performance Measurement: A Preliminary Review	
Z. Mustafa ¹ , N. R. M. Suradi ¹ , W. R. Ismail ¹ , K. A. M. Ali ¹ , Z. M. Ali ¹ , M. Abdullah ¹ , F. Idris ² , M. R. Ab Hamid ³	572
Simple Version of the Linear Business Cycle Model Anton Abdulbasah Kamil ^{1*} and Adam Baharum ^{2*}	57/
	576
Modelling The Relationships Between US and Selected Asian Stock Markets ¹ Mohd Tahir Ismail*, ² Rosmanjawati Abdul Rahman	580
Diffusive Logistic Equations with Single Time Delay and Variable Time Impulses	5 00
Jalina Widjaja	586
Developing a Model for Estimating Emission Caused by Vehicles at a Junction of Main Street Agus Salim Harahap, Herman Mawengkang	589
The Pricing of European Type Asian Options with Geometric Averages Mila Novita ¹ , Erline Natalia ² , Novita Dwi Hapsari ³ , Widia Desrianti ⁴)	593
The Use of Genetic Algorithm for Selecting The Eigenvectors In Pca Method For Face Image Recognition Suryadi MT, Yudi Satria, Helmiyati, Rahmi Rusin	on 600
Model of Portfolio Insurance with Optimal Strike Price Novriana Sumarti*, Muhammad Syamsuddin, Rieske Hadianti	607
	007
Cheapest Insertion - Convex Hull Approach to Euclidean TSP Abdulah Fajar* ^l , Nanna Suryana Herman ^{#2} , Nur Azman Abu ^{#3}	611
A Dynamic Nelson-Siegel Model for Indonesian Government Bond Yield Rates <i>R. Rahmawati</i> * ^{1,2} , <i>M. Syamsuddin</i> ^{1,3} , <i>R. Hadianti</i> ^{1,4} , <i>S. Afriani</i> ^{1,5} , <i>F. Damayanti</i> ^{1,6} , <i>A.P. Wulandari</i> ^{1,7}	616
Generalized reduced gradient method For earthquake resistant of Foundations Abdul Hakam	623
Modeling Dependence of Claim Amount between Different Claim Types using Copula Yulia Resti ¹ , Noriszura Ismail ² & Saiful Hafizah Jaaman ²)	629
The Role of Mathematics to Determine <i>Kiblat</i> Direction <i>Akhsanul In'am</i>	633
Dialog of Features Characteristics by Using Venn Diagram for Object Detecti	
Rahmadi Kurnia	637
Comparison of Vector Stream Cipher Based on Modification of Chebyshev Polynomial Function Santi Indarjani ¹ , Bety Hayat Susanti ² and Juniati ³	644
Workforce planning problem with demand changes Rustam Effendi Pasaribu,Rusly Siagian,Sabar,Saprida Montaria, Zulkifli, Herman Mawengkang	649
On solving the Plant Cycle Location Problem Henry Nainggolan, Misnawati, Nurtaito Sianturi, Rizky Ismalinda, Roslinawati, Herman Mawengkang	662
A mixed integer linear programming model for multi-item inventory problems	
Toha, Khairunisa Siregar, Lin Risnawati, Lord Byron Silalahi, Marsito, Herman Mawengkang, Integer programming model for operational aircraft maintenance routing problem with side constraints	664
Suhardi ¹ , Simon Petrus Sebayang ² , Benar ³ , Ardianta ⁴ , and Herman Mawengkang ⁵	672
Developing Mathematical Models for Analyzing The Performance Of A Growing Team Yulidar ¹ , Zunaida Sitorus ² , Makmur Tarigan ³ , Satriawan Taruna ⁴ , Gim Tarigan ⁵ , and Herman Mawengk	680 kang

Discovering Relationship of Single Word Entities from Weblogs Taufik Fuadi Abidin, Rasudin Abubakar, Alim Misbullah, Jufri Wahyudi	694
Evaluation on Fitness Assignment Methods for Multi-objective Examination Timetabling Problems <i>Taufiq Abdul Gani</i> ^a , <i>Nurmaulidar</i> ^b , <i>Ahamad Tajudin Khader</i> ^c	699
Face Recognition Using Smooth Support Vector Machine Based On Eigenfaces ^{1,2*} Muhammad Furqan, ² Abdullah Embong, ² Suryanti Awang, ² Santi W. Purnami, ² Sajadin Sembiring	708
Mathematical Simulation Circuit for the Scale Model of Geometry Normalized Electromagnetic System (GNES) Equipment *Muhammad Syukri Surbakti ^{1,2} , *Mohd Zubir Mat Jafri ² , *Lim Hwee San ² , ^Norhaslinda Mohamed-Tahr	715 rin^2
Handling Missing Values in Multiclass Multisurface Proximal Support Vector Machines Taufik Edy Sutanto	718
Traffic Flow Simulation on Simple Continuum Model Sri Mardiyati ¹ and Helen Burhan ²	724
Prototyping of a Gradient-Based Edge Detection Algorithm Design Indra Yasri*	728
A Mixed Integer Linear Programming Model for Capacity Selection Problem in Logistics Networks Optimisation	722
Muhammad Izman Herdiansyah ¹	732
Modeling Contagiousness of Diarrhea Diseases: A Spatial Probit Model Yusep Suparman*, Imam Munandar Fajari, Gatot Riwi Setyanto	743
Model of Growth Population by Modification of Malthus and Gomperz Growth Models <i>Georgina M. Tinungki</i>	746
Stability Analysis and Maximum Profit of Wangersky-Cunningham Population Model with Time Delay	and
Constant Effort of Harvesting Syamsuddin Toaha	751
Fuzzy correlation to contruct interaction function in group decision making Marwan Harahap	761
A Tabu Search with EST-SPT Algorithm for the Job Shop Scheduling Problem Opim Salim Sitompul	767
Finite Volume Method Based Analysis of Gas Flow in Two-Stroke Engine <i>Tulus</i>	772
An Improved Direct Feasible Search Approach for Solving Mixed-Integer Non Linear Programming Problems	
Elly Rosmaini, Herman Mawengkang	777
Optimization Methods for The Second Order Multiresponse Surface Model of Mixture Designs $^{1}Ruslan$, $^{2}Susanti$ L, $^{2}Purhadi$, $^{2}Sony$ S	781
The Application Of Classification Tree Method To Determine The Profile of Indonesian People Based on The Factors Which Significantly Influence The Attitudes Toward Avian Influenza Rianti Setiadi	n 785
Binary Stuctural Equation Model And Its Application To Find The Relationship Between "Knowledge A The Spread Of Avian Influenza" And "Preventive Actions Taken Toward Avian Influenza" <i>Rianti Setiadi</i>	bout 791
Applications of Fuzzy Number Max-Plus Eigenvalues on Queuing Networks with Fuzzy Activity Times M. Andy Rudhito ^{1*} , Sri Wahyuni ² , Ari Suparwanto ³ , and F. Susilo ⁴	
Distribution-free Test for Stability of Run-off Triangle Aceng Komarudin Mutaqin ^{1,2} , Dumaria Rulina Tampubolon ² , Sutawanir Darwis ²	802

The Optimization of Second Step of Crude Palm Oil Transesterification with Methanol Using Potas	sium
Hydroxide Catalyst Tirena Bahnur Siregar* and Sutarman	809
A Decision Analysis Model in Sustainable Land Revitalization Planning Using Participatory Approach Rahmawati Pane, Herman Mawengkang	813
Developing a Direct Search Algorithm For Solving the Capacitated Open Vehicle Routing Problem Tarno, Abdillah, Sudarman, Esmina Simatupan, Loide Naiborhu, Seprianti Harahap, Herman Mawengkang	819
An Improved Feasible Neighbourhood Search Approach For Solving The Capacitated Facility Location Problem	
Pramana, Tiopan Rahmat Siregar, Adil Pangaribuan, Sudarman Siringoringo, Indra Maryanti, Herman Mawengkang	826
Finding Shortest Path in Networks with Uncertain Arc Length Syamsul Qomar, Dewi Monalisa, Agus Budianto, Rosmartina, Lasma Nurhaida Silitonga, Yulis Purwaningsih, Herman Mawengkang	832
Developing a Direct Search Algorithm For Solving the Capacitated Open Vehicle Routing Problem Sudarman, Tarno, Abdillah, Esmina Simatupang, Loide Naiborhu, Seprianti Harahap, Herman Mawengkang	837
Development of Solow Growth Model Hamidah Nasution ¹ and Herman Mawengkang ² Matematika MIPA UNIMED, ²	844
Deciding Hub Location in a Communication Network Ummi Habibah, Rosimanidar, Purnawanto, Herman Mawengkang	848
A Constrained Based Approach for Handling The Multi-Period Single-Sourcing Problem Afnaria, Herman Mawengkang 860	
A Goal Programming Approach for a Class of Possibilistic Portfolio Selection Model Rina Filia Sari, and Herman Mawengkang	873
Optimization Model for Land Management Problems under Uncertainty Almira Amir	882
An Optimization Model for Irrigation Water Distribution Networks Alfred Hasiholan Silalahi,Rahmanan Dalimunthe,Isabella Bangun, Tiramah Simanjuntak,Surya Ningsih,Herman Mawengkang	888
Developing a constrained search approach for solving systems of nonlinear equations Pasukat Sembiring	893
On The Sufficient Condition for Solvability of Infinite Horizon LQ Problem Subject to DAE Systems <i>Muhafzan</i> ¹	902
Experimental Modelling of Domestic Tourist In West Sumatera by Using Separate Spline Function Susila Bahri	907
Structural Equation Approach for Non-Normality Data: With Reference to Modeling of Health Index Ferra Yanuar ^{1*} , Kamarulzaman Ibrahim ² , Abdul Aziz Jemain ³	909
Size Ramsey Number for P_3 and T_3 Des Welyyanti	916
Arranging The Examination Schedule By Using Graph Coloring Algorithm ¹ Budi Rahmadya, ² Narwen	918
Monitoring Mangrove Rehabilitation In Tsunami Affected Area Using High Resolution Satellite Images <i>Muzailin Affan</i>	926
Numerical Solution of Iron Corrosion Problem Based on Condensation Chemical Property ¹ Arid Fatahillah, ² Basuki Widodo	929
Constructing Model for Survival Data by the Makeham-like Unproportional Hazard Adhitya Ronnie Effendie ^{1*} , Danardono ² , Subanar ³	934

Comparative Analysis of Fuzzy Logic and Multiple Linear Regression On Forecasting of Brand Switching Based on the Value of Consumer Dissatisfaction, the Characteristics of Product Categories And Explore to Needs of Variation	_			
	939			
The Implementation Of Finite Difference Method To Simulation Of Transient Temperature Distribution Of Three-Dimensional Heat Conduction In Cartesian Coordinate System	Of 948			
Hufri				
Power Series Solution Of Non-Linear First Order Differential Equation Systems Zulakmal	954			
Model Reduction Using LMIs Jenizon	958			
Identification Algorithm For Temporal Logics Yahma Wisnani	960			
An edge consecutive edge magic total labeling on some classes of tree Kiki A. Sugeng and Denny R. Silaban	966			
On (a,d)-vertex antimagic labeling of circulant graphs Bong N. Herawati and Kiki A. Sugeng	970			
Determination of The Sex of Hawksbill Sea Turtle (Eretmochelys imbricata) by Using Logistic Regression Hazmira Yozza, Hilda Yohana, Izzati Rahmi HG, Kurniadi Ilham	n 974			

Distribution-free Test for Stability of Run-off Triangle

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Abstract

The estimation of outstanding claims liability (OCL) in long-tail insurance business is often based on run-off triangles. There are several methods to estimate OCL based on run-off triangles which assume a stable pattern of payments by development period within each accident period. These methods are run-off techniques. The chain ladder method is probably the most popular one of the run-off techniques. It is important to test the assumption of the stability of the payments pattern. This paper offers a distribution-free approach to test that stability. The approach is motivated by the basic chain ladder assumption and two additional assumptions in the paper by Mack (1993). In this paper, all link ratios in each development are tested if they are not statistically different from its development factor. For each development period, if at least one of the link ratios is statistically different from its development factor, then we conclude that the payment pattern is unstable. If at least one of the development periods is not stable, then we conclude that the run-off triangle is unstable. The critical values of the test are computed using simulation. Two run-off triangle data are used to illustrate the approach.

Keywords: chain ladder method, development factor, link ratio, outstanding claims liability, run-off triangle.

1. Introduction

In general insurance business, an insurance company must estimate its liability to make future payments for claims that have arisen on or before the valuation date. This liability is called outstanding claims liability (OCL). An insurance company needs to estimate OCL because there is delay between the occurrence of claim and the time of claim is settled.

Statistical methods are usually used to estimate OCL because of the uncertainty of OCL, since the amount and timing of future claims are unknown. The degree of uncertainty depends on whether the insurance business is short-tail or long-tail business (Olofsson, 2006). A short-tail insurance business is when the delay between the occurrence of a claim and the settlement is short, often less than a year, such as automobile physical damage insurance, fire insurance, and earthquake insurance. A long-tail business is an insurance business where the delay between the occurrence of a claim and the settlement is long, more than a year, such as third party liability, marine insurance, reinsurance, medical malpractice, and worker compensation.

It is common to estimate OCL for long-tail business based on run-off triangle data (for example, see Mack, 1993; England, and Verrall, 2002; Verrall, 2002; Panning, 2004; de Jong, 2006; Pinheiro et al., 2006; and Olofsson, 2006). Run-off triangle data provides a summary of underlying data set with individual claim figures (Antonio et al., 2006). There are several methods to estimate OCL based on run-off triangle data that assume a stable pattern of payments by development period within each accident period. These methods called run-off techniques (Hart et al., 1996). The chain ladder method is probably the most popular one of the run-off techniques for estimating OCL. The main reason for this is its simplicity and the fact that it is distribution-free (Mack, 1993). It is frequently used as a gold standard (benchmark) because of its generalized use and easy to apply (de Alba, 2006).

It is important to test the assumption of stability of the payments pattern. If we use run-off techniques to estimate OCL based on unstable run-off data, then the estimation result may be inappropriate. This paper offers a distribution-free approach to test that stability. This approach is motivated by the basic chain ladder assumption and two additional assumptions in the paper by Mack (1993). In this paper, all link ratios in each development are tested if they are statistically equal to its development factor or not. For each development period, if at least one of the link ratios is not statistically equal to its development factor, then we conclude that the payment pattern is unstable. If at least one of the development periods is not stable, then we conclude that the run-off triangle is unstable. The critical values of the test are computed using simulation.

The remainder of the paper is structured as follows. Section 2 presents the chain ladder method. The stability test is discussed in Section 3. The algorithm to compute the critical values of the test using simulation is given in Section 4. Section 5 provides two numerical examples to illustrate the approach.

2. The Chain Ladder Method

Let D_{ij} represent the incremental claim amounts (or the number of claims) in the development period j (j = 1, (2, ..., n), corresponding to the accident period i (i = 1, 2, ..., n). The cumulative claims are defined by

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$$C_{ij} = \sum_{t=1}^{j} D_{it}$$
; $i = 1, 2, \dots, n$.

We consider C_{ij} are random variables of which we have observations if i + j = 2, 3, ..., n + 1 (run-off triangle). The pattern of the available data can be seen in Table 1.

Table 1Run-off Triangle Data

Accident							
Period	1	2		j	• • •	n-1	n
1	C_{11}	C_{12}		C_{1j}	• • •	$C_{1,n-1}$	C_{1n}
2	C_{21}	C_{22}		C_{2j}	• • •	$C_{2,n-1}$	
:	÷	:	:	:			
i	C_{i1}	C_{i2}		C_{ij}			
:	:	:					
n-1	$C_{n-1,1}$	$C_{n-1,2}$					
n	$C_{n,1}$						

From now on, and without loss of generality, we consider that C_{ij} are the cumulative claims amounts. The aim of the run-off techniques is to estimate all the empty cells in Table 1 to obtain the estimate of OCL. The most popular one of the run-off techniques for estimating OCL is the chain ladder method (Mack, 1993). The chain ladder method has three assumptions as follows.

- (1) There are development factors $\lambda_1, \lambda_2, \cdots, \lambda_{n-1} > 0$ with expectation $E(C_{i,j+1}|C_{i1}, \cdots, C_{ij}) = C_{ij}\lambda_j, i = 1, 2, \dots, n; j = 1, 2, \dots, n 1.$
- (2) $\{C_{i1}, \dots, C_{in}\}, \{C_{k1}, \dots, C_{kn}\}, i \neq k$, are independent.
- (3) Variance $Var(C_{i,j+1}|C_{i1},\cdots,C_{ij}) = C_{ij}\sigma_j^2, i = 1,2,...,n; j = 1,2,...,n-1,$ with unknown parameters $\sigma_j^2, j = 1,2,...,n-1$.

The first assumption is the basic chain ladder assumption and the last two assumptions are the additional assumptions stated by Mack (1993). The estimates of the development factors are

$$\hat{\lambda}_{j} = \frac{\sum_{i=1}^{n-j} C_{i,j+1}}{\sum_{i=1}^{n-j} C_{i,j}}, \quad j = 1, 2, \dots, n-1,$$
(1)

and the estimates of parameters σ_i^2 , j = 1, 2, ..., n - 2 are

$$\hat{\sigma}_{j}^{2} = \frac{1}{n-j-1} \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_{j} \right)^{2}, j = 1, 2, ..., n-2.$$
(2)

Mack (1993) suggested that if $\hat{\lambda}_{n-1} = 1$ and if the claims development is believed to be finished after n-1 years, then $\hat{\sigma}_{n-1} = 0$, otherwise

$$\hat{\sigma}_{n-1}^2 = min\left(\frac{\hat{\sigma}_{n-2}^4}{\hat{\sigma}_{n-3}^2}, min(\hat{\sigma}_{n-3}^2, \hat{\sigma}_{n-2}^2)\right). \tag{3}$$

Theorem 1: Let $D = \{C_{ij} | i + j = 2, ..., n + 1\}$ be the data set are observed. Under the assumptions (1) and (2) we have

$$E(C_{in}|D) = C_{i,n+1-i}\lambda_{n+1-i} \cdot \dots \cdot \lambda_{n-1}.$$

Theorem 2: Under the assumptions (1) and (2) the estimators $\hat{\lambda}_j$, j = 1, 2, ..., n - 1, are unbiased and uncorrelated.

Theorems 1 and 2 have been proven by Mack (1993), and are used to estimate OCL in the chain ladder method.

Theorem 3: Under the assumptions (1), (2), and (3) the estimators $\hat{\sigma}_i^2$, j = 1, 2, ..., n - 2, are unbiased.

Proof: Begin with the following result, which will be used later:

$$\sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_{j} \right)^{2} = \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} + \lambda_{j} - \hat{\lambda}_{j} \right)^{2} \\
= \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} \right)^{2} + 2 \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} \right) (\lambda_{j} - \hat{\lambda}_{j}) + \sum_{i=1}^{n-j} C_{ij} (\lambda_{j} - \hat{\lambda}_{j})^{2} \\
= \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} \right)^{2} + 2(\lambda_{j} - \hat{\lambda}_{j}) \sum_{i=1}^{n-j} (C_{i,j+1} - C_{ij} \lambda_{j}) + (\lambda_{j} - \hat{\lambda}_{j})^{2} \sum_{i=1}^{n-j} C_{ij} \\
= \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} \right)^{2} + 2(\lambda_{j} - \hat{\lambda}_{j}) (\hat{\lambda}_{j} - \lambda_{j}) \sum_{i=1}^{n-j} C_{ij} + (\lambda_{j} - \hat{\lambda}_{j})^{2} \sum_{i=1}^{n-j} C_{ij} \\
= \sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j} \right)^{2} - (\hat{\lambda}_{j} - \lambda_{j})^{2} \sum_{i=1}^{n-j} C_{ij} . \tag{*}$$

We also have (from independence of $\{C_{i1}, \dots, C_{in}\}, \{C_{k1}, \dots, C_{kn}\}, i \neq k\}$

$$Var(\hat{\lambda}_{j}) = Var\left(\frac{\sum_{i=1}^{n-j} C_{i,j+1}}{\sum_{i=1}^{n-j} C_{i,j}}\right) = \frac{1}{\left(\sum_{i=1}^{n-j} C_{i,j}\right)^{2}} \sum_{i=1}^{n-j} Var(C_{i,j+1}) = \frac{\sigma_{j}^{2}}{\left(\sum_{i=1}^{n-j} C_{i,j}\right)^{2}} \sum_{i=1}^{n-j} C_{i,j} = \frac{\sigma_{j}^{2}}{\sum_{i=1}^{n-j} C_{i,j}}$$

Take expectation in (*) to obtain

$$E\left[\sum_{i=1}^{n-j} C_{ij} \left(\frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_{j}\right)^{2}\right] = \sum_{i=1}^{n-j} C_{ij} E\left[\left(\frac{C_{i,j+1}}{C_{ij}} - \lambda_{j}\right)^{2}\right] - \sum_{i=1}^{n-j} C_{ij} E\left[\left(\hat{\lambda}_{j} - \lambda_{j}\right)^{2}\right]$$

$$= \sum_{i=1}^{n-j} C_{ij} Var\left(\frac{C_{i,j+1}}{C_{ij}}\right) - \sum_{i=1}^{n-j} C_{ij} Var(\hat{\lambda}_{j}) = \sum_{i=1}^{n-j} C_{ij} \frac{1}{\left(C_{ij}\right)^{2}} C_{ij} \sigma_{j}^{2} - \sum_{i=1}^{n-j} C_{ij} \frac{\sigma_{j}^{2}}{\sum_{i=1}^{n-j} C_{ij}}$$

$$= (n-j-1)\sigma_{i}^{2}.$$

Dividing both sides by (n-j-1) demonstrates that $\hat{\sigma}_i^2$ is an unbiased estimator of σ_i^2 .

3. The Stability Test

Run-off techniques to estimate OCL assume a stable pattern of payments by development period within each accident period. This section discusses a distribution-free approach to test that stability. This approach is motivated by the basic chain ladder assumption and two additional assumptions in the paper by Mack (1993).

The test will be based on the following theorem.

Theorem 4: Let

$$W_{ij} = \frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_j, \quad j = 1, 2, ..., n-1, i = 1, ..., n-j$$

Under the assumptions (1), (2), and (3) we have expectations and variances

$$E(W_{ij} | C_{i1}, \dots, C_{ij}) = 0,$$

$$Var(W_{ij} | C_{i1}, \dots, C_{ij}) = \left(\frac{1}{C_{ij}} - \frac{1}{\sum_{b=1}^{n-j} C_{bi}}\right) \sigma_j^2.$$

Proof: We use the abbreviations

$$E_{ij}(X) = E(X|C_{i1}, \dots, C_{ij}),$$

$$Var_{ij}(X) = Var(X|C_{i1}, \dots, C_{ij}).$$

For the expectations

$$E_{ij}\left(W_{ij}\right) = E_{ij}\left(\frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_j\right) = \frac{E_{ij}\left(C_{i,j+1}\right)}{C_{ij}} - E_{ij}\left(\hat{\lambda}_j\right) = \frac{C_{ij}\lambda_j}{C_{ij}} - \lambda_j = 0.$$

For the variances

$$\begin{aligned} Var_{ij}\left(W_{ij}\right) &= Var_{ij}\left(\frac{C_{i,j+1}}{C_{ij}} - \hat{\lambda}_{j}\right) = Var_{ij}\left(\frac{C_{i,j+1}}{C_{ij}} - \frac{\sum_{k=1}^{n-j} C_{k,j+1}}{\sum_{k=1}^{n-j} C_{kj}}\right) \\ &= Var_{ij}\left(\frac{C_{i,j+1}}{C_{ij}} - \frac{C_{i,j+1}}{\sum_{k=1}^{n-j} C_{kj}} - \frac{\sum_{k=1}^{n-j} C_{k,j+1}}{\sum_{k=1}^{n-j} C_{kj}}\right) \\ &= Var_{ij}\left(\left[\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right] C_{i,j+1} - \frac{\sum_{k=1}^{n-j} C_{k,j+1}}{\sum_{k=1}^{n-j} C_{kj}}\right) \\ &= Var_{ij}\left(\left[\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right] C_{i,j+1}\right) + Var_{ij}\left(\frac{\sum_{k=1}^{n-j} C_{k,j+1}}{\sum_{k=1}^{n-j} C_{kj}}\right) \\ &= \left[\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right]^{2} Var_{ij}\left(C_{i,j+1}\right) + \frac{1}{\left(\sum_{k=1}^{n-j} C_{kj}\right)^{2}} \sum_{k=1}^{n-j} Var_{ij}\left(C_{k,j+1}\right) \\ &= \left[\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right]^{2} C_{ij} \sigma_{j}^{2} + \frac{1}{\left(\sum_{k=1}^{n-j} C_{kj}\right)^{2}} \sum_{k=1}^{n-j} C_{kj} \sigma_{j}^{2} \\ &= \left(\frac{1}{C_{ij}} + \frac{C_{ij}}{\left(\sum_{k=1}^{n-j} C_{kj}\right)^{2}} - \frac{2}{\sum_{k=1}^{n-j} C_{kj}} + \frac{\sum_{k=1}^{n-j} C_{kj}}{\left(\sum_{k=1}^{n-j} C_{kj}\right)^{2}} \right) \sigma_{j}^{2} = \left(\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right) \sigma_{j}^{2}. \end{aligned}$$

Define the link ratios

$$R_{ij} = \frac{C_{i,j+1}}{C_{ii}}, \quad j = 1, 2, \dots, n-1, i = 1, \dots, n-j.$$
(4)

Run-off triangle is stable if all payments patterns in every development are stable. In other words, Run-off triangle is stable if all link ratios in every development are stable. For a given development j (j = 1, 2, ..., n - 1), all link ratios are stable if all the difference between link ratio R_{ij} and λ_j are small with high probability or all the difference between link ratio R_{ij} and $\hat{\lambda}_j$ are small relative to standard deviation of W_{ij} , $\sigma_{W_{ij}}$ with high probability. In statistical term, this means that probability

$$P(|R_{ij} - \hat{\lambda}_j| \le \varepsilon_j \sigma_{W_{ij}}) \ge p$$

or

$$P\left(\left|\frac{W_{ij}}{\sigma_{W_{ij}}}\right| \le \varepsilon_j\right) \ge p.$$

Because W_{ij} has a continuous distribution, the " \geq " sign in the above equation may be replaced by an "=" sign,

$$P\left(\left|\frac{W_{ij}}{\sigma_{W_{ij}}}\right| \le \varepsilon_j\right) = p. \tag{5}$$

Equation (5) will be used to test the stability of the run-off triangle. The detailed steps of the stability test of the run-off triangle are as follow

(1) All link ratios in each development period are tested if they are not statistically different from its development factor. For development j (j = 1, ..., n - 1), the test statistic is

$$\frac{W_{ij}}{\sigma_{W_{ij}}}, \quad j = 1, \dots, n - 1, i = 1, \dots, n - j,$$
(6)

where

$$W_{ij} = \frac{C_{i,j+1}}{C_{ii}} - \hat{\lambda}_j, \quad j = 1, 2, ..., n-1, i = 1, ..., n-j,$$
(7)

 $\sigma_{W_{ii}}$ is estimated by

$$\hat{\sigma}_{W_{ij}} = \sqrt{\left(\frac{1}{C_{ij}} - \frac{1}{\sum_{k=1}^{n-j} C_{kj}}\right) \hat{\sigma}_j^2}, \quad j = 1, 2, \dots, n-1, i = 1, \dots, n-j.$$
(8)

Algorithm to compute the critical values of the test, i.e. ε_j , is discussed in the next section. If the absolute value of the test statistic is less than or equal to the critical value, then we conclude that the link ratio is not statistically different from its development factor.

- (2) For each development period, if at least one of the link ratios is statistically different from its development factor, then we conclude that the payment pattern is unstable.
- (3) If at least one of the development periods is not stable, then we conclude that the run-off triangle is unstable.

4. The Critical Value of the Test

This section discusses the computation of the critical values of the stability test of run-off triangle simulation. The algorithm to compute critical value for development j is given as follows

- (1) Compute the estimated development factor $\hat{\lambda}_i$ using Equation (1).
- (2) Compute the estimated variance component $\hat{\sigma}_i^2$ using Equation (2) and (3).
- (3) Let C_{ij} $\{i=1,\cdots,n+1-j\}$ represent the cumulative claims in development j. Compute the estimated mean and the estimated variance of $C_{i,j+1}$ $\{i=1,\cdots,n-j\}$. The estimated means are $C_{ij} \hat{\lambda}_j$, $i=1,\ldots,n-j$, and the estimated variances are $C_{ij} \hat{\sigma}_i^2$, $i=1,\ldots,n-j$.
- (4) Generate the cumulative claims $C_{i,j+1}$ $\{i=1,\cdots,n-j\}$ from a distribution (for example lognormal, gamma, or Weibull) with mean $C_{ij}\hat{\lambda}_j$, $i=1,\ldots,n-j$ and variance $C_{ij}\hat{\sigma}_j^2$, $i=1,\ldots,n-j$.
- (5) Based on $C_{i,j+1}$ $\{i=1,\cdots,n-j\}$ obtained from Step (4) and $C_{i,j}$ $\{i=1,\cdots,n+1-j\}$ from the run-off triangle data, compute
 - (5.1) The estimated development factor $\hat{\lambda}_i^*$ using Equation (1),
 - (5.2) The estimated variance component $\hat{\sigma}_i^{2*}$ using Equation (2) and (3),
 - (5.3) W_{ij} ; i = 1, ..., n j, using Equation (7).
 - (5.4) $\hat{\sigma}_{W_{ii}}$; i = 1, ..., n j, using Equation (8).
 - (5.5) The absolute values of $W_{ij}/\hat{\sigma}_{W_{ij}}$.
- (6) Repeat Step (4) and (5) *N* times.
- (7) For a given ε ($\varepsilon = 1, \dots, 2.5$; with increment 0.01), compute $S_i = \#\{|W_{ij}/\hat{\sigma}_{W_{ij}}| \le \varepsilon\}/N, i = 1, \dots, n-j$
- (8) Compute mean of S_i , i = 1, ..., n j,

$$S = \frac{\sum_{i=1}^{n-j} S_i}{n-j}$$

- (9) Repeat Step (4) until Step (8) *M* times, then compute mean of *S*.
- (10) The critical value of the test is ε such that $S = 1 \alpha$, where α is the level of significance for the test.

5. The Numerical Examples

This section discusses the implementation of the stability test of run-off triangle for two data sets. The data set 1 are taken from Mack (1993) which were used by Verrall (1990). The data set 1 in Table 2 contain cumulative claims. The data set 2 are taken from de Jong (2006) in Table 3 relates to Automatic Facultative General Liability from the historical Loss Development Study. This triangle, called the AFG data that contain cumulative incurred claim amounts.

 Table 2

 Data Set 1 - Run-off Triangle Data (Cumulative Figures)

i	C_{i1}	C_{i2}	C_{i3}	C_{i4}	C_{i5}	C_{i6}	C_{i7}	C_{i8}	C_{i9}	C_{i10}
1	357848	1124788	1735330	2218270	2745596	3319994	3466336	3606286	3833515	3901463
2	352118	1236139	2170033	3353322	3799067	4120063	4647867	4914039	5339085	
3	290507	1292306	2218525	3235179	3985995	4132918	4628910	4909315		
4	310608	1418858	2195047	3757447	4029929	4381982	4588268			
5	443160	1136350	2128333	2897821	3402672	3873311				
6	396132	1333217	2180715	2985752	3691712					
7	440832	1288463	2419861	3483130						

8	359480	1421128	2864498				
9	376686	1363294					
10	344014						

Table 3Data Set 2 – AFG Data in Cumulative

i	C_{i1}	C_{i2}	C_{i3}	C_{i4}	C_{i5}	C_{i6}	C_{i7}	C_{i8}	C_{i9}	C_{i10}
1	5012	8269	10907	11805	13539	16181	18009	18608	18662	18834
2	106	4285	5396	10666	13782	15599	15496	16169	16704	
3	3410	8992	13873	16141	18735	22214	22863	23466		
4	5655	11555	15766	21266	23425	26083	27067			
5	1092	9565	15836	22169	25955	26180				
6	1513	6445	11702	12935	15852					
7	557	4020	10946	12314						
8	1351	6947	13112							
9	3133	5395								
10	2063	·	·	·	·					

Table 4 and 5 contains the test statistics and the critical values for data set 1 and data set 2, respectively. The run-off triangle for data set 1 is stable because in Table 4, all link ratios are not statistically different from its development factor (for each development, all test statistics is less than its critical value). However, the run-off triangle for data set 2 is unstable because in Table 5, the development periods 1, 2, 3, and 5 are not stable.

Table 4Test Statistics and Critical Values for Data Set 1

Test Statisties and Critical Values for Data Set 1										
,	Development									
l l	1	2	3	4	5	6	7	8		
1	0,183	0,144	0,131	0,049	0,085	0,036	0,011	0,011		
2	0,010	0,006	0,073	0,040	0,019	0,041	0,004	0,011		
3	0,448	0,023	0,001	0,055	0,068	0,033	0,007			
4	0,523	0,161	0,213	0,106	0,017	0,040				
5	0,550	0,089	0,079	0,000	0,032					
6	0,070	0,086	0,074	0,057						
7	0,336	0,099	0,016							
8	0,244	0,215								
9	0,070									
Critical	2,03	2,1	2,08	2,02	1,9	1,7	> 1,41	> 0,99		
Values	2,03	2,1	2,00	2,02	1,7	1,7	<i>-</i> 1,41	~ U,99		

Table 5Test Statistics and Critical Values for Data Set 2

_		1 Cot Diac	istics and	offical va.	uco for Du	ita Det 2			
i	Development								
ı	1	2	3	4	5	6	7	8	
1	0,652	0,896	0,803	0,368	0,940	1,583	0,000	1,000	
2	2,313*	0,743	2,038*	1,688	0,214	1,057	1,282	1,000	
3	0,138	0,249	0,526	0,195	1,010	0,349	1,166		
4	0,500	0,931	0,413	1,484	0,001	0,130			
5	1,169	0,103	0,685	0,019	1,809*				
6	0,304	0,490	0,734	0,842					
7	0,604	2,167*	0,622						
8	0,487	0,703							
9	0,463				•				
Critical Values	2,28	1,9	1,87	1,82	1,76	1,65	1,41	> 1,01	

^{*} means significance at $\alpha = 5\%$.

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